COMP2610/COMP6261 - Information Theory Lecture 5: Useful Discrete Probability Distributions

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August 5th, 2014

- Examples of application of Bayes' rule
 - Formalizing problems in language of probability
 - Eating hamburgers, detecting terrorists, document classification
- Frequentist vs Bayesian probabilities

Bayesian Inference

Bayesian inference provides us with a a mathematical framework explaining how to change our (prior) beliefs in the light of new evidence.



Prior: Belief that someone is sick

Likelihood: Probability of testing positive given you are sick

Posterior: Probability of being sick given you test positive

Mark Reid and Aditya Menon (ANU) COMP2610/COMP6261 - Information Theory

- The Bernoulli and binomial distribution
- Estimating probabilities from data
- Bayesian inference for parameter estimation



- The Bernoulli Distribution
- 2 The Binomial Distribution
- 3 Parameter Estimation
- 4 Bayesian Parameter Estimation





2 The Binomial Distribution

3 Parameter Estimation

4 Bayesian Parameter Estimation



Consider a binary variable $X \in \{0, 1\}$. It could represent many things:

- Whether a coin lands heads or tails
- The presence/absence of a word in a document
- A transmitted bit in a message
- The success of a medical trial

Often, these outcomes (0 or 1) are not equally likely

What is a general way to model such an X?

The variable X takes on the outcomes

$$X = egin{cases} 1 & ext{ probability } heta \ 0 & ext{ probability } 1 - heta \end{cases}$$

Here, 0 $\leq \theta \leq$ 1 is a parameter representing the probability of success

For higher values of θ , it is more likely to see 1 than 0

• e.g. a biased coin

The Bernoulli Distribution

By definition,

$$p(X = 1| heta) = heta$$

 $p(X = 0| heta) = 1 - heta$

More succinctly,

$$p(X = x|\theta) = \theta^{x}(1-\theta)^{1-x}$$

The Bernoulli Distribution Definition

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More succinctly,

$$p(X = x|\theta) = \theta^{x}(1-\theta)^{1-x}$$

This is known as a Bernoulli distribution over binary outcomes:

$$p(X = x|\theta) = \text{Bern}(x|\theta) = \theta^{x}(1-\theta)^{1-x}$$

Note the use of the conditioning symbol for θ ; will revisit later

The expected value (or mean) is given by:

$$egin{aligned} \mathbb{E}[X| heta] &= \sum_{x\in\{0,1\}} x\cdot p(x| heta) \ &= 1\cdot p(X=1| heta) + 0\cdot p(X=0| heta) \ &= heta. \end{aligned}$$

The variance (or squared standard deviation) is given by:

$$\mathbb{V}[X|\theta] = \mathbb{E}[(X - \mathbb{E}[X])^2] = \theta(1 - \theta).$$

Example: Binary Symmetric Channel

Suppose a sender transmits messages s that are sequences of bits

The receiver sees the bit sequence (message) t

Due to noise in the channel, the message is flipped with probability $0 \le f \le 1$



We can think of r as the outcome of a random variable, with conditional distribution given by:



If *E* denotes whether an error occurred, clearly

$$p(E = e) = Bern(e|f).$$



2 The Binomial Distribution

- 3 Parameter Estimation
- 4 Bayesian Parameter Estimation
- 5 Wrapping up

Suppose we perform N independent Bernoulli trials

- e.g. we toss a coin N times
- e.g. we transmit a sequence of N bits across a noisy channel

Each trial has probability $\boldsymbol{\theta}$ of success

What is the distribution of the number of times (m) that X = 1?

- e.g. the number of times we obtained *m* heads
- e.g. the number of errors in the transmitted sequence

The Binomial Distribution Definition

Let

$$Y = \sum_{n=1}^{N} X_n$$

where $X_n \sim \text{Bern}(\theta)$

Then Y has a binomial distribution with parameters N, θ :

$$p(Y = m) = Bin(m|N, \theta) = {N \choose m} \theta^m (1 - \theta)^{N-m}$$

where

$$\binom{N}{m} = \frac{N!}{(N-m)!m!}$$

is the # of ways we can we obtain m heads out of N coin flips

It is easy to show that:

$$\mathbb{E}[Y] = \sum_{m=0}^{N} m \cdot \operatorname{Bin}(m|N,\theta) = N\theta$$
$$\mathbb{V}[Y] = \sum_{m=0}^{N} (m - \mathbb{E}[m])^2 \cdot \operatorname{Bin}(m|N,\theta) = N\theta(1-\theta)$$

• Follows from linearity of mean and variance

- Ashton is an excellent off spinner. The probability of him getting a wicket during a cricket match is $\frac{1}{4}$.
- His coach, Darren, commands him to get 10 wickets in a particular game.
 - What is the probability that he will get exactly three wickets?
 - What is the expected number of wickets he will get?
 - What is the probability that he will get at least one wicket?

The Binomial Distribution:

Example: Distribution of the Number of Wickets



Figure: Histogram of the binomial distribution with N=10 and $\theta=0.25$. From Bishop (PRML, 2006)

1 The Bernoulli Distribution

2 The Binomial Distribution

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Consider the set of observations $\mathcal{D} = \{x_1, \dots, x_N\}$ with $x_i \in \{0, 1\}$:

- The outcomes of a sequence of coin flips
- Whether or not there are errors in a transmitted bit string

Each observation is the outcome of a random variable X, with distribution

$$p(X = x) = \text{Bern}(x|\theta) = \theta^{x}(1-\theta)^{1-x}$$

for some parameter θ

We know that

$$X \sim \text{Bern}(x|\theta) = \theta^x (1-\theta)^{1-x}$$

But often, we don't know what the value of θ is

- The probability of a coin toss resulting in heads
- The probability of the word *defence* appearing in a document about sports

What would be a reasonable estimate for θ from \mathcal{D} ?

Say that we observe

$$\mathcal{D} = \{0, 0, 0, 1, 0, 0, 1, 0, 0, 0\}$$

Intuitively, which seems more plausible: $\theta = \frac{1}{2}$? $\theta = \frac{1}{5}$?

Say that we observe

$$\mathcal{D} = \{0, 0, 0, 1, 0, 0, 1, 0, 0, 0\}$$

If it were true that $\theta = \frac{1}{2}$, then the probability of this sequence would be

$$p(\mathcal{D}|\theta) = \prod_{i=1}^{10} p(x_i|\theta)$$
$$= \prod_{i=1}^{10} \frac{1}{2}$$
$$= \frac{1}{2^{10}}$$
$$\approx 0.001$$

Say that we observe

$$\mathcal{D} = \{0, 0, 0, 1, 0, 0, 1, 0, 0, 0\}$$

If it were true that $\theta = \frac{1}{5}$, then the probability of this sequence would be

$$egin{aligned} p(\mathcal{D}| heta) &= \prod_{i=1}^{10} p(x_i| heta) \ &= \left(rac{1}{5}
ight)^2 \cdot \left(rac{4}{5}
ight)^8 \ &pprox 0.007. \end{aligned}$$

We can write down how likely ${\mathcal D}$ is under the Bernoulli model. Assuming independent observations:

$$p(\mathcal{D}|\theta) = \prod_{i=1}^{N} p(x_i|\theta) = \prod_{i=1}^{N} \theta^{x_i} (1-\theta)^{1-x_i}$$

We call $L(\theta) = p(\mathcal{D}|\theta)$ the likelihood function

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Maximum likelihood principle: We want to maximize this function wrt θ

The parameter for which the observed sequence has the highest probability

Maximising $p(\mathcal{D}|\theta)$ is equivalent to maximising $\mathcal{L}(\theta) = \log p(\mathcal{D}|\theta)$

$$\mathcal{L}(heta) = \log p(\mathcal{D}| heta) = \sum_{i=1}^N \log p(x_i| heta) = \sum_{i=1}^N [x_i \log heta + (1-x_i) \log(1- heta)]$$

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Setting $\frac{d\mathcal{L}}{d\theta} = 0$ we obtain:

$$heta_{\mathsf{ML}} = rac{1}{N} \sum_{i=1}^{N} x_i$$

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The proportion of times x = 1 in the dataset \mathcal{D} !

The Bernoulli Distribution:

Parameter Estimation — Issues with Maximum Likelihood

Consider the following scenarios:

- After N = 3 coin flips we obtained 3 'tails'
 - What is the estimate of the probability of a coin flip resulting in 'heads'?
- In a small set of documents about sports, the words *defence* never appeared.
 - What are the consequences when predicting whether a document is about sports (using Bayes' rule)?

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These issues are usually referred to as overfitting

- Need to "smooth" out our parameter estimates
- Alternatively, we can do Bayesian inference by considering priors over the parameters

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The Bernoulli Distribution:

Parameter Estimation: Bayesian Inference

Recall:



If we treat θ as a random variable, we may have some prior belief $p(\theta)$ about its value

• e.g. we believe θ is probably close to 0.5

Our prior on θ quantifies what we believe θ is likely to be, before looking at the data

Our posterior on θ quantifies what we believe θ is likely to be, after looking at the data

The likelihood of X given θ is

$$\operatorname{Bern}(x|\theta) = \theta^{x}(1-\theta)^{1-x}$$

For the prior, it is mathematically convenient to express it as a Beta distribution:

$$\mathsf{Beta}(\theta|a,b) = \frac{1}{Z(a,b)} \theta^{a-1} (1-\theta)^{b-1},$$

where Z(a, b) is a suitable normaliser

We can tune a, b to reflect our belief in the range of likely values of θ

Beta Prior Examples



Beta Prior and Binomial Likelihood:

Beta Posterior Distribution

Recall that for $\mathcal{D} = \{x_1, \dots, x_N\}$, the likelihood under a Bernoulli model is:

$$p(\mathcal{D}|\theta) = \theta^m (1-\theta)^\ell$$

where $m = \sharp(x = 1)$ and $\ell \stackrel{\text{\tiny def}}{=} N - m = \sharp(x = 0)$.

Beta Prior and Binomial Likelihood:

Beta Posterior Distribution

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For the prior $p(\theta|a, b) = \text{Beta}(\theta|a, b)$ we can obtain the posterior:

$$p(\theta|\mathcal{D}, a, b) = \frac{p(\mathcal{D}|\theta)p(\theta|a, b)}{p(\mathcal{D}|a, b)}$$
$$= \frac{p(\mathcal{D}|\theta)p(\theta|a, b)}{\int_0^1 p(\mathcal{D}|\theta)p(\theta|a, b)d\theta}$$
$$= \text{Beta}(\theta|m+a, \ell+b).$$

Beta Prior and Binomial Likelihood:

Beta Posterior Distribution

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Can use this as our new prior if we see more data!

Beta Posterior Distribution

Now suppose we choose θ_{MAP} to maximise $p(\theta|\mathcal{D})$

One can show that

$$heta_{\mathsf{MAP}} = rac{m+a-1}{N+a+b-2}$$

c.f. the estimate that did not use any prior,

$$\theta_{ML} = \frac{m}{N}$$

The prior parameters a and b can be seen as adding some "fake" trials!

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- Distributions involving binary random variables
 - Bernoulli distribution
 - Binomial distribution
- Bayesian inference: Full posterior on the parameters
 - Beta prior and binomial likelihood \rightarrow Beta posterior
- Reading: Mackay §23.1 and §23.5; Bishop §2.1 and §2.2

• The entropy and its properties